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Good Luck or Good Policy? An Expectational Theory of Macro-Volatility Switches

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Résumé: Nous considérons un modèle où les agents sont segmentés sur un petit nombre d'îlots d'information en fonction du signal qu'ils reçoivent sur les anticipations des autres et qui sinon aurait un équilibre unique. Même si les agents observent parfaitement les fondamentaux, des équilibres d'exubérance rationnelle (REX) peuvent survenir lorsqu'ils mettent du poids sur les signaux des anticipations pour affiner leurs prévisions. Un apprentissage adaptatif perpétuel peut déclencher des sauts entre l'équilibre où seuls les fondamentaux sont pris en compte et un REX. Cela engendre des changements dans la volatilité macroéconomique sans pour autant que la politique monétaire ou que la distribution des chocs exogènes changent dans le temps. Dans ce contexte, une politique de ciblage de l'inflation peut réduire la complémentarité des anticipations et prévenir l'exubérance rationnelle, et ceci bien que son effet soit non-monotone.

Classification JEL: E3, E5, D8.

Mots-clés: volatilité non-fondamentale; apprentissage perpétuel; co-variations des anticipations; prévisionnistes professionnels.

Abstract: In an otherwise unique-equilibrium model, agents are segmented into a few informational islands according to the signal they receive about others' expectations. Even if agents perfectly observe fundamentals, rational-exuberance equilibria (REX) can arise as they put weight on expectational signals to refine their forecasts. Constant-gain adaptive learning can trigger jumps between the equilibrium where only fundamentals are weighted and a REX. This determines regime switching in macro volatility despite unchanged monetary policy and time-invariant distribution of exogenous shocks. In this context, a tight inflation-targeting policy can lower expectational complementarity preventing rational exuberance, although its effect is non-monotone.

JEL Classification: E3, E5, D8.

Keywords: non-fundamental volatility; perpetual learning; comovements in expectations; professional forecasters.

1 Introduction

What are the determinants of switches in the volatility of macro-variables? In principle, a persistent reduction in the amplitude of business fluctuations can be thought to be either the result of *good policy*, namely a change of policy by some major actor within the economy, or of *good luck*, that is, a decrease of volatility of the exogenous shocks hitting the economy. It is not always easy to distinguish between the two. An example is provided by the intense debate on the sources of the "great moderation" in the 80s' (Stock and Watson, 2002; McConnell M.M. and Perez-Quiros, 2000).

This paper presents a simple model where the introduction of signals about expectations of others jointly with adaptive learning can generate shifts in macro volatility with unchanged monetary policy and time-invariant distribution of exogenous shocks. Still, it assigns to monetary policy an important but ambiguous role. Policies of tight targeting on inflation can in fact prevent the possibility of regimes of high volatility, although marginal hardening is counter productive once high volatility occurs.

I consider a monopolistic competition economy where producers have to set their price before knowing the aggregate price. A policy maker enforces a flexible targeting rule according to his preferred trade-off between output gap and inflation volatility. In this model the actual output gap responds to actual inflation that in turns responds to the producers' average expectation about current inflation. Under homogenous information a unique rational expectation equilibrium exists. In this context two main twists are introduced.

First, the economy is split into two symmetrical islands. On each island, the expectation of each producer is his own noisy perception of the forecast of an island-specific type of professional forecaster. The latter is intended as a medium or a statistical office that releases reports on the future course of inflation. Thus, there is an information transmission channel that maps professional forecasts into naive producers' expectations depending on the distribution of the perception noises across the population.

The professional forecasters perfectly observe all the fundamental determinants of inflation, but they also receive a private signal of the other professional forecaster's expectation. That is, each professional forecaster can anticipate the forecasts of the other with some uncertainty. Expectational signals are the only sources of heterogeneity between the professional forecasters since each one observes a private signal from the other's expectation.

The introduction of heterogeneous expectational signals can give origin

to a multiplicity of rational expectation equilibria. A fundamental equilibrium always exists in which experts just use fundamental information and do not put weight on expectational signals. Then, two rational exuberance equilibria can arise in which experts put weight on expectational signals self-fulfilling rational exuberance. In particular, a multiplicity of equilibria exist under two conditions: i) the monetary policy is not aggressive enough about the inflation target and ii) the map from professional forecasts to producers' expectations entails an amplification of the non-fundamental component. The sum of these two effects can provide the degree of complementarity needed to self-fulfil the predictive power of expectational signals. In the case of a rational exuberance equilibrium, non-fundamental volatility driven by observational noises transmits to actual inflation entailing a regime of higher volatility.

As in a typical sunspot equilibrium, at a rational exuberance equilibrium it is optimal to put weight on some non-fundamental signal if everybody does the same. Nevertheless rational exuberance equilibria and sunspot equilibria¹ are essentially different. The latter require that a commonly understood exogenous signal drives the coordination of agents' beliefs, the former instead originate with *heterogeneous* signals that are *endogenous* to the forecasting rule. Expectational signals are not simple coordination devices, but they entail a signal extraction problem that sustains a multiplicity of equilibria. In fact, as with the model at hand, rational exuberance equilibria can exist where typical sunspots do not.

The second twist is to explore the consequences of professional forecasters acting like econometricians, that is using linear regressions on observables to form their forecasts (Evans and Honkapohja, 2001). In particular, I explore the possibility that they learn with a constant gain, so that exponentially decreasing weights are given to earlier data. This class of learning algorithms is particularly suited to learn about stochastic processes that are potentially open to sudden structural changes.

The paper proves that whenever rational exuberance equilibria exist at least one of them is learnable under adaptive learning². Moreover, the learnability of rational exuberance equilibria and of the fundamental equilibrium coexists in a large region of the parameter space. Therefore, in this region, constant gain learning selects among them and potentially triggers unpre-

¹For comprehensive reviews on sunspots see Benhabib and Farmer (1999) and Guesnerie (2001).

²This is an other difference with typical sunspot equilibria that instead are learnable only in some limited cases under special representations (see Evans and McGough (2011)).

dictable and endogenous jumps among learnable equilibria following several lucky or unlucky expectational aggregate shocks. This is possible as long as a small number of types is considered, so that, when expectational signals are weighted, the impact of observational noises does not vanish into the aggregation. In this way different regime switching in volatility can occur despite unchanged monetary policy and a time-invariant distribution of exogenous shocks.

Nevertheless, the monetary authority still has an important role. By implementing a flexible targeting rule the central bank can amplify or dampen the impact of aggregate expectation on actual inflation. In particular, a sufficiently high focus on price stability can prevent a multiplicity of equilibria. However, only a *marginal* hardening of the monetary policy can increase inflation volatility when rational exuberance is already in play. The change of monetary policy must be drastic to be beneficial. Only a gradual focus on price stability is fated to generate periods of even higher volatility. In this sense, *good policy* is not strictly necessary, but it is sufficient to prevent *bad luck* if appropriately conducted.

Finally, in Section 6, the dynamic system is simulated for few parametrizations in the cases of two and more-than-two informational islands. With a finite number of islands the same qualitative results obtain, although the quantitative dimension becomes less important as the number of islands increases.

2 Related literature

Angeletos and Werning (2006) and Hellwig and others (2006) have emphasized the importance of endogenous signals in restoring multiplicity in the static version of the benchmark currency attack model when agents are privately uncertain about the fundamentals. As shown in Gaballo (2012), endogenous signals do not just restore multiplicity, but they can be a source of a multiplicity in otherwise unique-equilibrium models.

The attempt to reconcile macro-volatility regime switches with constant gain learning is shared with Branch and Evans (2007). In their model, stochastic volatility is the result of an evolutionary competition among different misspecified predictors. Sargent and others (2008) also use misspecified constant-gain learning to explain the rise and fall of South American inflation. The results in this paper do not rely on misspecified forecasting rules, but rather on informational frictions.

Other works have proved that a very aggressive monetary policy stabi-

lizes macro-volatility when agents learn. Orphanides and Williams (2005) show that when agents use perpetual adaptive learning schemes then reduced price fluctuations are reflected in a low volatility of expectations, which in turn stabilizes output. In Branch and others (2009) agents choose a level of attention by balancing the effect of active monetary policy on output volatility. Adam (2009) proves the local uniqueness of an equilibrium where fully rational inattentive firms can more easily process data when prices are stable, contributing to overall stability. In the present model there are two important differences: the change in monetary policy has to be drastic to be beneficial when exuberance is already in play, and *good policy* is not necessary for the stability of the system.

3 Baseline model

3.1 Aggregate supply with uninformed suppliers

This section introduces a simple model to address the issue of regime switching in volatility driven by heterogeneous expectations. The basic framework, derived in appendix, is a textbook monopolistic competition model populated by a continuum of suppliers with unit mass along the lines of Woodford (2003).

To allow for heterogeneous expectations, I assume firms price their product before knowing the aggregate price and the aggregate output. Their log-linearized optimal pricing rule is

$$p_{i,t} = E_{t-1}^i p_t + \omega_y E_{t-1}^i y_t + \omega_c c_{i,t} - \tilde{z}_{t-1}, \quad (1)$$

where, $E_{t-1}^i(\cdot)$ denotes the conditional expectation operator of producer i , p_t is the aggregate price, y_t is the output gap, $c_{i,t}$ is the consumption by workers type i , \tilde{z}_{t-1} is a predetermined technology shock drawn from a normal distribution centred on zero with finite variance σ_z , finally ω_y and ω_c are deep parameters. In other words, each firm knows the consumption of its own workers, but they set price before knowing others' current pricing and supply.³ After aggregation we have

$$p_t = E_{t-1} p_t + \omega_y E_{t-1} y_t + \omega_c y_t - \tilde{z}_{t-1} \quad (2)$$

where $E_{t-1}(\cdot) \equiv \int E_{t-1}^i(\cdot) di$ denotes the average expectation across uninformed firms. Hence, the aggregate supply (AS) is written as

$$y_t = \omega_c^{-1} (p_t - E_{t-1} p_t) - \zeta E_{t-1} y_t + (1 + \zeta) z_{t-1}, \quad (3)$$

³This informational assumption implies a weaker departure from full knowledge than the one originally postulated in Woodford (2003).

where $\zeta = \omega_y/\omega_c$ is the CES coefficient of labor disutility and $z_{t-1} \equiv \tilde{z}_{t-1}/(1 + \zeta)\omega_c$ is the rescaled technology shock. In particular $\omega_c^{-1} = (1 + \theta\zeta) > 0$ captures the degree of strategic complementarity in price setting which increases in the degree of market power θ and in the elasticity of labor disutility ζ . The relation above is a kind of new classical Phillips curve encompassing that in Lucas (1973), and Kydland and Prescott (1977), extended to incorporate the effect of the average expectation about the aggregate supply.

3.2 Monetary policy

A monetary authority has the instruments to successfully implement the following flexible targeting rule

$$\pi_t + \phi y_t = \pi^* + \tilde{\eta}_t, \quad (4)$$

where $\pi_t \equiv p_t - p_{t-1}$ is the inflation rate, π^* is the announced inflation target, $\tilde{\eta}_t$ are i.i.d. white noise shocks with finite variance interpreted as monetary transmission frictions and $\phi \geq 0$ represents the degree of flexibility of the targeting rule. Specifically, $\phi = 0$ entails the most restrictive monetary regime with actual inflation being on average equal to the target. As ϕ increases, the response to inflation becomes weaker during recessions and stricter in expansionary periods. This specification is a simple and general way to embody the implications of different degrees of policy-maker tolerance to deviations from the inflation target conditional on output deviations from the steady state.

The structure of the targeting rule is known. Therefore each supplier's expectation on the aggregate supply is pinned down by her expectation on the aggregate price according to $E_{t-1}^i y_t = \phi^{-1} \pi^* - \phi^{-1} E_{t-1}^i \pi_t$ where I assume the monetary disturbance is in fact truly unpredictable, that is $E_{t-1}^i \tilde{\eta}_t = 0$. At the aggregate level this relation implies

$$E_{t-1} y_t = \phi^{-1} \pi^* - \phi^{-1} E_{t-1} \pi_t \quad (5)$$

that is, the average suppliers' belief on output gap $E_{t-1} y_t$ is mapped one-to-one to the average suppliers' belief on current inflation $E_{t-1} \pi_t$.

3.3 The equilibrium under homogeneous expectations

Plugging (4) and (5) to substitute respectively y_t and $E_{t-1} y_t$ into (3), current inflation is given by the following reduced form:

$$\pi_t = \boldsymbol{\alpha}' \mathbf{z}_{t-1} + \beta E_{t-1} \pi_t + \eta_t, \quad (6)$$

where

$$\boldsymbol{\alpha} \equiv \frac{(1 + \zeta)\omega_c}{\phi + \omega_c} \begin{bmatrix} 1 \\ 1 \end{bmatrix}, \quad \mathbf{z}_{t-1} \equiv \begin{bmatrix} \pi^* \\ -\phi z_{t-1} \end{bmatrix}, \quad \beta \equiv \frac{\phi - \zeta\omega_c}{\phi + \omega_c}, \quad \eta_t \equiv \frac{\omega_c}{\phi + \omega_c} \tilde{\eta}_t,$$

which describes the course of actual inflation given producers' expectations. Notice β measures the impact of the average expectations of current inflation. In particular, β lies in $(-\zeta, 1]$ and monotonically decreases in ϕ , the degree of tightness of the monetary policy. Notice moreover that $\boldsymbol{\alpha}'/(1 - \beta) = [1 \ 1]$. It is straightforward to find the rational expectation equilibrium in the case of homogenous expectations.

Definition 1 *The minimal state variable equilibrium of the model (MSV) is characterized by*

$$\begin{aligned} \pi_t &= \pi^* - \phi z_{t-1} + \eta_t, \\ y_t &= z_{t-1} + \omega_c^{-1} \eta_t \\ E_{t-1} \pi_t &= \bar{\pi}_t \equiv \pi^* - \phi z_{t-1} \end{aligned}$$

that is, the unique stationary sequence of inflation rates, output gaps and individual expectations that satisfies (1), (3) and (6) under the restriction of homogeneous expectations, namely $E_{t-1}^i \pi_t = E_{t-1} \pi_t$ for each i .

The minimal state variable equilibrium arises when producers expects the fundamental inflation rate $\bar{\pi}_t$, that is, the one predicted by the fundamentals of the economy: the inflation target and the predetermined technology shock. The mere existence of the MSV cannot explain macro-volatility switches unless there are structural changes in the distribution of the exogenous shocks or changes in the policy of the monetary authority.

4 Introducing heterogeneous expectations

In the following I introduce heterogeneity in expectations. I develop the simplest (and more transparent) case when uninformed producers are symmetrically segmented into two informational islands, namely i and j . The extension to a finite number of islands is analytically cumbersome, but it can be easily obtained numerically. Some simulations are provided and discussed at the end of the paper.

Professional Forecasters. On each island there is an island-specific type of professional forecasters whose only aim is to truthfully provide the best projections of the inflation course (in the mean square error sense) to

uninformed producers inhabiting their own island. Professional forecasters neither produce nor consume. They can be thought of as media or statistical offices that serve a certain industrial district, as well as institutional agencies that release reports on the expected course of the economy based on available information.

The information set of the professional forecaster type ι is

$$\Omega_{t-1}^\iota \equiv [\{\mathbf{z}_\tau\}_{\tau=-\infty}^{t-1}, \{s_{\iota,\tau}\}_{\tau=-\infty}^{t-1}], \quad (7)$$

with

$$s_{\iota,t-1} \equiv E_{t-1}^j \pi_t + \eta_{\iota,t-1},$$

where $\eta_{\iota,t-1}$ is a type-specific white noise measurement error drawn from a normal distribution $N(0, \sigma)$ with zero mean and finite variance σ , and $E_{t-1}^j \pi_t$ is the price expectation of the professional forecaster type j . The noises in expectational signals are orthogonal to fundamental variables and independently distributed in time. The information set of professional forecaster type j is a mirror image. Professional forecasters observe fundamentals (the inflation target and the predetermined technology shock) and a signal of the simultaneous expectation of the professional forecaster on the other island. Expectational signals capture uncertainty about others' expectations measured by the size of σ . The case of perfect information obtains in the limit of $\sigma \rightarrow 0$, whereas the case of no information about others' forecasts arises in the limit of $\sigma \rightarrow \infty$.

Individual-specific noises in the expectational signals are the only source of heterogeneity in the information set of the professional forecasters. This feature allows us to divide their forecasting problem into two sequential tasks: estimating the fundamental rate of inflation and estimating non-fundamental fluctuations from the fundamental. In fact, the professional forecasters form expectations about the fundamental rate $\bar{\pi}_t$ conditioning on the commonly observed fundamental variables according to

$$E_{t-1}^\iota \bar{\pi}_t = E_{t-1}^j \bar{\pi}_t = \bar{\pi}_t^e \equiv \mathbf{a}' \mathbf{z}_{t-1}, \quad (8)$$

where \mathbf{a} is a bidimensional weight to be determined in equilibrium. The expected fundamental rate is the same since both types of professional forecasters use the same fundamental information.

Even in the case that the professional forecasters correctly estimate the fundamental inflation rate, unexpected fluctuations are caused by the exogenous unobserved disturbance η_t . However, the professional forecasters are uncertain whether such departures are truly exogenous or partly due to

expectations of non-fundamental fluctuations held on the other island. In this case the expectational signal would be a useful predictor of course of the actual inflation. Therefore, the professional forecasters estimate non-fundamental fluctuations around the estimated fundamental rate using the linear rule⁴

$$\begin{aligned} E_{t-1}^i \pi_t - \bar{\pi}_t^e &= b \left(E_{t-1}^j \pi_t + \eta_{i,t-1} - \bar{\pi}_t^e \right), \\ E_{t-1}^j \pi_t - \bar{\pi}_t^e &= c \left(E_{t-1}^i \pi_t + \eta_{j,t-1} - \bar{\pi}_t^e \right), \end{aligned}$$

where b and c are weights to be determined in equilibrium. Following this rule, each professional forecaster expects a departure of the actual inflation from the fundamental rate proportional to his own noisy observation of the other's mirror-like expectation. Solving the equations above for the professional forecasts, we can rewrite them as functions of the expected fundamental rate and observational errors. We have

$$E_{t-1}^i \pi_t = \bar{\pi}_t^e + \frac{bc}{1-bc} \eta_{j,t-1} + \frac{b}{1-bc} \eta_{i,t-1}, \quad (9a)$$

$$E_{t-1}^j \pi_t = \bar{\pi}_t^e + \frac{bc}{1-bc} \eta_{i,t-1} + \frac{c}{1-bc} \eta_{j,t-1}, \quad (9b)$$

with $bc \neq 1$. Notice that both types of professional forecasts are determined by both b and c . In particular they collapse to (8) if and only if both b and c are equal to zero.

Producers. Actual inflation reacts to the aggregate expectation of uninformed agents, and not directly to that of professional forecasters. I assume that producers do not have a particular theory of how the economy works, but rather they rely on the expectations of a more sophisticated agent. Imagine that although the reports of the professional forecasters are public they can be misperceived or interpreted in different ways. For the sake of simplicity, I assume that the expectation of producer i on the island ι ,

$$E_{t-1}^i \pi_t = E_{t-1}^i \pi_t + \eta_{i,t-1}, \quad (10)$$

deviates from the one of the professional forecaster type i by a white-noise measurement error $\eta_{i,t-1}$ which encapsulates the outcome of a truly private stochastic process of decoding. In other words, producers' expectations are

⁴Notice that by this assumption I restrict my analysis to linear rational expectation equilibria, that is, equilibria in which deviations from a steady state are linear combinations of the shocks. This is a typical choice in the literature as not much is known about the tractability of non-linear rational expectation equilibria.

just noisy understandings of the professional forecasts on their own island. The average expectation across producers is therefore equal to

$$E_{t-1}\pi_t = \frac{1}{2} (E_{t-1}^i \pi_t + E_{t-1}^j \pi_t) + \int \eta_{i,t-1} di. \quad (11)$$

The relation above entails a map from the average professional forecast to the average expectation across producers. The properties of this map depends on the aggregation of perception noises.

The transmission channel. This structure shapes a relation between experts and the private sector, which is illustrated in figure 1.

[figure 1 about here]

Two types of professional forecasters equally affect the average expectation calculated over a continuum of suppliers. The average expectation yields the actual inflation rate as implied by (6). The professional forecasters observe fundamentals and have noisy observations of each others' expectations.

Notice that the last term in (11) shapes the degree of neutrality of the transmission channel from experts to the private sector: it is zero if and only if the cross-sectional correlation of the perception errors is zero. Here instead I want to allow perception noises to be correlated with the non-fundamental component of the professional forecast they rely on. In other words, I want to account for the possibility that the emphasis in the reports about the non-fundamental nature of expected fluctuations could systematically bias the way producers read these reports. To capture this effect one can express $\eta_{i,t} \sim N(0, \sigma_{i,t})$ in the following way:

$$\eta_{i,t-1} = \gamma(E_{t-1}^i \pi_t - \bar{\pi}_t^e) + \epsilon_{i,t}, \quad (12)$$

where $\epsilon_{i,t}$ is a i.i.d. shock normally distributed according to $N(0, \sigma_\epsilon)$ across agents and time. With this form, the average expectation can be written simply as

$$E_{t-1}\pi_t = \bar{\pi}_t^e + (1 + \gamma) \left(\frac{E_{t-1}^i \pi_t + E_{t-1}^j \pi_t}{2} - \bar{\pi}_t^e \right), \quad (13)$$

where γ measures the effect of the transmission channel from professional to naive forecasts. Notice that the impact of the non-fundamental component of the professional forecasts on the average expectation is amplified or dampened according to the sign and size of γ .

This structural assumption, as others in the macro literature (notably Calvo pricing or some types of matching functions), is a modeling device that enables to replicate with a parsimonious parameterization well-documented empirical regularities at the aggregate level. In fact, the relations (8), (9) and (13) together are able to capture stylized facts recovered from the examination of expectation surveys. Mankiw, Reis and Wolfers (2004) prove important claims about the relation between disagreement in the Survey of Professional Forecasters (SPF) and the Michigan Survey of Consumers' expectations: i) both cross-sectional distributions are closely centred on the right value of inflation, that is, average expectations are almost rational expectations; ii) both present similar slow reactions to news about fundamental macroeconomic data that only account for a small fraction of the overall volatility⁵; iii) both types of expectations present substantial correlation, so they seem to react to similar fundamental *and* non-fundamental shocks. On top of that, the observation that the expectations of the private sector are systematically more volatile than professional ones, although they exhibit a strong correlation but a similar slow reaction to fundamental news⁶, supports a positive value of γ .⁷ That is, the private sector is more sensitive than professional forecasters to non-fundamental shocks, which in this model are captured by the noises in expectational signals.

Moreover, this hypothesis allows a much more direct understanding of the dynamics. In fact, β and γ will be the only parameters determining the expectational complementarity between the professional forecasters' expectations, which is ultimately what matters for the emergence of a multiplicity of equilibria. With different assumptions on the correlation of expectational signals one can still obtain a multiplicity without relying on any exogenous amplification channel, but at the cost of much more involved conditions. For example, one possibility is to assume a correlation between fundamentals and expectational signals as shown in Gaballo (2012). The reader interested in more microfounded mechanisms is referred to that work. The assumption

⁵In particular, this can reflect the fact that expectations adaptively incorporates news concerning fundamental data.

⁶Carrol (2003) presents an epidemiological model where the private sector's expectations exhibit even slower reactions to news of macroeconomic fundamentals than do those of professional forecasters.

⁷Using (9a), (10) and (12) we can write

$$E_{t-1}^i \pi_t = \bar{\pi}_t^e + (1 + \gamma) \left[\frac{bc}{1 - bc} \eta_{j,t-1} + \frac{b}{1 - bc} \eta_{i,t-1} \right] + \epsilon_{i,t}.$$

so that $\text{Var}(E^i \pi_t - \bar{\pi}_t^e) = (1 + \gamma)^2 \text{Var}(E^i \pi_t - \bar{\pi}_t^e) + \sigma_\epsilon$ while both distributions are centred on $\bar{\pi}_t^e$ across time.

made here focuses attention on the central aim of the current paper, which is to show how constant gain learning can trigger endogenous regime switching in macro volatility with unchanged monetary policy while keeping the distribution of exogenous shocks time-invariant.

5 Multiple Learnable Equilibria

5.1 Rational Expectations Equilibria

It is possible now to recover the actual law of motion (ALM) as a function of exogenous shocks only, parameterized by the weights of the forecasting rules used by the two types of professional forecasters. Plugging (9) and (13) into (6), the ALM for inflation is

$$\pi_t = \boldsymbol{\alpha}' \mathbf{z}_{t-1} + \beta \bar{\pi}_t^e + \frac{\beta^*}{2} \left(\frac{b(1+c)}{1-bc} \eta_{i,t-1} + \frac{c(b+1)}{1-bc} \eta_{j,t-1} \right) + \eta_t \quad (14)$$

where $\beta^* \equiv \beta(1+\gamma)$ measures the final impact of professional forecasts of non-fundamental fluctuations. Equation (14) features *temporary equilibria* of the model, that is, it describes the course of actual inflation for *given* coefficients (\mathbf{a}, b, c) of the forecasting rules (8)-(9). The rational expectation equilibrium values of (\mathbf{a}, b, c) imply instead that such forecasting rules are optimal linear projections of available information, so that the professional forecasters' mistakes are orthogonal to available information collected in (7).⁸

Definition 2 *The projected T-map (or simply T-map) is a map $T(\mathbf{a}, b, c) : \mathbb{R}^4 \rightarrow \mathbb{R}^4$ giving, for any parametrization (\mathbf{a}, b, c) of the inflation process (14), the weights $(T_{\mathbf{a}}, T_b, T_c)$ such that*

$$\begin{aligned} \mathbf{E}[\mathbf{z}_{t-1} (\pi_t - T_{\mathbf{a}} \mathbf{z}_{t-1})] &= 0, \\ \mathbf{E}[(E_{t-1}^j \pi_t + \eta_{i,t-1} - \bar{\pi}_t^e) (\pi_t - \bar{\pi}_t^e - T_b (E_{t-1}^j \pi_t + \eta_{i,t-1} - \bar{\pi}_t^e))] &= 0, \\ \mathbf{E}[(E_{t-1}^i \pi_t + \eta_{j,t-1} - \bar{\pi}_t^e) (\pi_t - \bar{\pi}_t^e - T_c (E_{t-1}^i \pi_t + \eta_{j,t-1} - \bar{\pi}_t^e))] &= 0, \end{aligned}$$

that is, the available information is orthogonal to the forecast errors.

⁸When shocks are normally distributed, this restriction is enough to pin down the forecasting rule that identifies the correct conditional distribution of the variable to be forecasted.

One can work out the orthogonal restrictions above to obtain the explicit form for the T-map

$$\begin{aligned} T_{\mathbf{a}}(\mathbf{a}) &= \boldsymbol{\alpha} + \beta \mathbf{a}, \\ T_b(b, c) &= \frac{\beta^*}{2} \left(\frac{b(1+c)(1+c\rho) + c(1+b)(c+\rho)}{1+c^2+2c\rho} \right), \\ T_c(b, c) &= \frac{\beta^*}{2} \left(\frac{b(1+c)(b+\rho) + c(1+b)(1+b\rho)}{1+b^2+2b\rho} \right), \end{aligned}$$

where $\rho \equiv \text{corr}[\eta_{i,t}, \eta_{j,t}]$ is the correlation among the simultaneous observational errors of the professional forecasters. Notice that the T-map does not depend on the variance σ of observational errors.

Definition 3 *Rational expectation equilibria are a sequence of actual inflation rates, output gaps and individual expectations as defined by (3)-(9)-(14) for a triple $(\mathbf{a}, b, c) = T(\mathbf{a}, b, c)$ that is a fix point of the T-map.*

The following proposition states the existence of multiple determinate equilibria.

Proposition 4 *The system has one or three rational expectation equilibria. The minimal state variable equilibrium (MSV) with $(\mathbf{a}, b, c) = (1, 1, 0, 0)$ always exists. A high rational exuberance equilibrium (hREX) with $(\mathbf{a}, b, c) = (1, 1, b_+, c_+)$, and a low rational exuberance equilibrium (lREX) with $(\mathbf{a}, b, c) = (1, 1, b_-, c_-)$, where*

$$b_{\pm} = c_{\pm} = \frac{\beta^* - (2 - \beta^*)\rho \pm 2\sqrt{(\beta^* - 1)(1 - \rho^2)}}{2 - \beta^*(1 + \rho)},$$

both exist provided $\beta^* > 1$.

Proof. In the appendix. ■

The signal extraction problem entailed by the expectational signals is crucial to the existence of REX. This occurs whenever $|\text{corr}[s_{i,t}, s_{j,t}]| \neq 1$, that is the correlation between the expectational signals is not perfect. In particular, notice that the correlation between expectational signals is endogenous to the given weights b and c . At the REX values $b_{\pm} = c_{\pm}$ is

$$\text{corr}[s_{i,t}, s_{j,t}] = \frac{2b_{\pm} + (1 + b_{\pm}^2)\rho}{2\rho b_{\pm} + 1 + b_{\pm}^2},$$

whose absolute value is one only at $b_{\pm} = c_{\pm} = 1$ obtained for $\beta^* = 1$ or $\rho = \pm 1$. These are the limit points at which heterogeneity of the information

sets vanishes. In such a case expectations are homogeneous and so the MSV is the only equilibrium. For all the other cases in which $b_{\pm} = c_{\pm}$ exist and are different from one - that is whenever $\beta^* > 1$ - two REX exist.

Figure 2 plots T_b for four different calibrations. Given the symmetric nature of the problem, REX lie at the intersections of T_b with the bisector. Line *a*) is obtained for $\beta^* = 0.8$ and $\rho = 0$. In this case there is a unique intersection at the MSV values $b = c = 0$. For β^* greater than one, two rational exuberance equilibria (REX) emerge. Ceteris paribus, increasing values of ρ widen the distance between the low and the high REX values (contrast *b*) with *c*). Finally, an extreme parametrization such as that entailing line *d*) gives rise to negative low REX values.

[figure 2 about here]

The intuition for the result is simple. A multiplicity of equilibria is due to the non-linearity of the T-map originated by the endogeneity and heterogeneity of the expectational signals. These two features create externalities to the individual forecasting problem and REX as coordination failures. In fact, both types of professional forecasters could obtain a lower variance of their forecast errors if both coordinated on the MSV values. But suppose type i puts weight on his expectational signal. Now $E_{t-1}^i \pi_t$ reacts to $\eta_{i,t-1}$, which therefore affects the course of inflation π_t . Then $s_{j,t-1}$ is now partly informative of actual inflation fluctuations away from the estimated fundamental rate so that type j wants to condition on it, coming in full circle. In other words, if type i deviates from the fundamental rate then this creates an incentive for type j to do the same. When the expectational complementarity is strong enough ($\beta^* > 1$) then this mechanism is self-reinforcing and a multiplicity of determinate equilibria - consisting of two REX and a MSV - exists; otherwise putting zero weight on the signals is the only equilibrium forecasting rule.

5.2 Adaptive learning

This section explores the learnability of equilibria, and in particular the possibility of the professional forecasters being stuck in a REX when they learn adaptively with a constant-gain algorithm. The concept of learnability refers to the nature - stable or unstable - of the learning dynamics around the equilibria. Consider that $(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ are recursively estimated with constant

stochastic gradient (CSG) according to

$$\mathbf{a}_t = \mathbf{a}_{t-1} + g \mathbf{z}_{t-1} (\pi_t - \mathbf{a}'_{t-1} \mathbf{z}_{t-1}), \quad (15a)$$

$$b_t = b_{t-1} + g (E_{t-1}^j \pi_t + \eta_{i,t-1} - \bar{\pi}_t^e) (\pi_t - E_{t-1}^i \pi_t), \quad (15b)$$

$$c_t = c_{t-1} + g (E_{t-1}^i \pi_t + \eta_{j,t-1} - \bar{\pi}_t^e) (\pi_t - E_{t-1}^j \pi_t), \quad (15c)$$

where g is a fixed gain, $(\pi_t - E_{t-1}^i \pi_t)$ is the forecast error and $(E_{t-1}^j \pi_t + \eta_{i,t-1} - \bar{\pi}_t^e)$ is the noisy observed displacements of others' expectations from the estimated fundamental rate.⁹ If both estimates b_t and c_t are close to zero, the professional forecasters discard non-fundamental information and forecast the fundamental rate given by (15a). If this is not the case, considering noisy observations actually improves the accuracy of forecasts and the system can possibly converge on a learnable REX.

Definition 5 *An equilibrium entailed by $(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ is locally learnable under a constant-gain learning algorithm if and only if there exists a sufficiently small gain \tilde{g} and some neighborhood $\mathfrak{S}(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ of $(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ such that for each initial condition $(\mathbf{a}_0, b_0, c_0) \in \mathfrak{S}(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ and positive gain $g \leq \tilde{g}$ the estimates converge almost surely in distribution to the equilibrium values.*

The particular constant-gain algorithm (15) is similar to the recursive version of OLS where the estimated correlation matrix is kept equal to the identity matrix and the gain is fixed¹⁰. CSG converges to an ergodic distribution centred on a fixed point of the T-map whenever recursive OLS asymptotically converges (to a point). CSG like any constant gain learning rule, exhibits perpetual learning since more weight is given to more recent data. This makes this class of algorithms particularly suitable for learning structural changes. In this model, CSG has the advantage of showing

⁹The self-referential nature of the problem implies that the coefficients that agents are estimating change over time during the convergence to the REE values. Therefore agents should estimate them using a Kalman-filtering technique. Nevertheless, adaptive learning literature a la Evans and Honkapohja (2001) (to which I refer to) disregards this aspect assuming agents are boundedly rational. On the robustness of the findings the interested reader is referred to McGough (2003).

¹⁰The CSG algorithm is derived as the optimal solution to the minimization of the forecast error variance provided agents are "sensitive" to risk in a particular form. For details see Evans, Honkapohja and Williams (2005). The CSG formula is obtained by setting the covariance matrix equal to the unitary matrix in the recursive expression for the constant gain OLS (cgOLS). In order to obtain adjustments comparable with cgOLS the gain has to be rescaled by the covariance matrix of the regressors. None of the theoretical result would change assuming cgOLS instead of CSG. Nevertheless, CSG performs better in the simulations as endogenous switches obtain more easily.

convergence to the equilibria and, at the same time, the possibility of endogenous and unpredictable shifts from the MSV to a REX. Recursive OLS on the contrary converge at the cost of a huge stickiness of the dynamics after relatively few repetitions.

To check learnability one needs to investigate the Jacobian of the T-map. If all eigenvalues of the Jacobian of the T-map computed at the equilibrium values have real parts less than one, then the equilibrium is stable under learning (Marcet and Sargent, 1989; Evans and Honkapohja, 2001). The Jacobian for the T-map takes the form

$$JT(\mathbf{a}, b, c) = \begin{pmatrix} \beta & 0 & 0 & 0 \\ 0 & \beta & 0 & 0 \\ 0 & 0 & \frac{dT_b(b,c)}{db} & \frac{dT_b(b,c)}{dc} \\ 0 & 0 & \frac{dT_c(b,c)}{db} & \frac{dT_c(b,c)}{dc} \end{pmatrix}$$

where the explicit forms of the partial derivatives are recovered in the appendix.

Figure 3 shows a numerical analysis for the whole parameter range spanned by β^* and ρ . Remember that a necessary condition for the learnability of equilibria is always $\beta < 1$, given that β is the largest eigenvalue associated with the learning dynamics of the fundamental component of actual inflation. This is assumed in the following discussion.

[figure 3 about here]

The MSV is the only learnable equilibrium in the region $\beta^* < 1$. In the white area a learnable high REX arises besides the MSV. In this case the learning algorithm selects between them. As ρ increases in modulus, MSV and hREX are both learnable for lower values of β^* . Specifically, as ρ approaches unity for a sufficiently high value of β^* the low REX becomes the unique learnable equilibrium. In contrast, as ρ decreases for sufficiently high value of β^* the system presents no learnable equilibria.

Necessary condition for arising and learnability of REX in this model is that $\gamma > \beta^{-1} - 1$, that is, the private sector commits perception errors that positively covary with the estimated departures of actual inflation from the fundamental rate. In that respect, the transmission channel in the economy plays an essential role in that it provides the degree of expectational complementarity needed for the emergence of REX.

Whenever learnable REX exist, the distances between the equilibria measured on the bisector in figure 2 are indicative of the size of the basins of attraction. In particular, at least in the range considered, the T-map behaves

like a cubic yielding three dynamic equilibria. As usual, the equilibrium in the middle is either unstable and works as the threshold between two basins of attraction or it is the only stable equilibrium with a basin of attraction lying between the other two. In the case of lines *b*) and *c*) in figure 2, the low REX divides the basins of attraction of the MSV and of the high REX. In particular, as β^* increases the latter enhances whereas the former shrinks.

Constant gain algorithms trigger continuous temporary escapes from the equilibrium values. Such escapes can displace the system from one basin of attraction to another. This can occur since only two (or a small number of) islands are considered so that expectational shocks do not vanish in the aggregation and affect current inflation. In this way regime switching in volatility can alternate despite unchanged monetary policy and time-invariant distribution of exogenous shocks.

5.3 Aggregate volatility and monetary policy

From a qualitative point of view non-fundamental volatility generated by rational exuberance has the features of demand-side volatility. The model implies that as long as a decrease in the overall variance is due to a shift from a REX to a MSV, output gaps and hours should exhibit a sudden decrease in volatility whereas labor productivity should be unaffected¹¹. This effect is in sharp contrast with a pure good luck theory for which all these variables should present an equal decrease in variation resulting from a weakening of technological shocks.

[figure 4 about here]

Figure 4 plots excess volatility at the learnable REX measured in observational error variance units. For β^* values close to unity excess volatility triggered by the high REX is huge but it initially decreases very soon and then slowly increases again as β^* increases. This turns out to be crucial in the evaluation of the monetary policy conduct. A more aggressive monetary policy, that is a lower ϕ , lowers β and so β^* . For values of ϕ such that $\beta^* < 1$ rational exuberance cannot emerge. Nonetheless, when rational exuberance is already in play, a gradual decrease in ϕ enhances volatility as β^* approaches unity. In other words, more focus on price stability is not locally optimal. Therefore, the monetary authority must implement a drastic change of policy to avoid high-volatility periods, whereas partial adjustments are harmful whenever rational exuberance is already in play.

¹¹Part of the evidence uncovered by Galí and Gambetti (2009) seems to confirm these predictions relative to the "great moderation" puzzle.

6 Simulations

Examples of endogenous and unpredictable regime switching in volatility are provided in this section. I choose calibrations to allow a simple comparison with the analytical results. I set $\pi^* = 0$ and $z_t = 0$ at each t so that the fundamental rate of inflation is equal to zero at all times. This requires a_t to be estimated as a constant converging to the equilibrium values of $\hat{a} = 0$. This choice is without loss of generality. The exogenous shocks are all Gaussian white noises with $\sigma = 0.01$ variance. The values of the benchmark parametrization are: $\phi = 6.5$; $\zeta = 3$; $\theta = 3$ so that β is about 0.94. In all the figures the following conventions hold. From top to bottom the three boxes in each figure respectively display the series for: actual inflation, the forecast of the fundamental rate, and the evolution of the weights on expectational signals. Whenever a multiplicity of equilibria exist, their values are indicated by dotted lines in the last two boxes.

6.1 Two islands

Figure 5 displays the benchmark case of convergence in distribution to the MSV values for $\rho = \gamma = 0$. The gain is settled at $g = 0.01/\sigma$. The factor σ has been included in the gain so that the adjustments of both b_t and c_t are substantially equal to those obtained with constant gain recursive least squares around the MSV for 0.01 (see footnote 7). Notice how constant-gain learning generates continuous small displacements away from the MSV values. Such displacements are temporary escapes and do not substantially affect the variance of actual inflation displayed in the first box.

[figure 5 about here]

[figure 6 about here]

Figure 6 shows the occurrence of an endogenous structural change affecting the volatility of actual inflation in a persistent and substantial way. The parametrization of figure 5 is modified only by setting $\gamma (\simeq 0.14)$ so that $\beta^* = 1.08$. Therefore the learnable high REX arises for $b = c = 1.78$ as shown by figure 3 in correspondence with curve b). For about 1300, periods the dynamics of figures 5 and 6 are indistinguishable. Nevertheless, as estimates approach the low REX values, the course of actual inflation changes dramatically even if the economy is perturbed by the same series of exogenous shocks. In particular, as estimates overcome the low REX values (this happens after about 2000 periods) the dynamics enters in the bigger basin of

attraction of the high REX¹². The extra non-fundamental volatility is about three times the MSV one. This change could be easily misidentified by an external observer as a truly exogenous increase in non-fundamental volatility. Notice also that expectational volatility affects the learning process of the estimated fundamental rate contributing to the overall volatility.

[figure 7 about here]

Figure 7 plots the perils of a timid monetary policy conduct. The economy is initially trapped in a high REX at $b = c = 2.33$ as β^* is set around 1.16 (and $\rho = 0$). At time 1000 a modest hardening of the monetary policy occurs such that β^* falls to 1.01. The transition to the new high REX equilibrium at $b = c = 1.35$ takes about 600 periods (the gain is $g = 0.01/\sigma$). Although monetary policy is now tighter, inflation volatility between time 1600 and time 2400 is higher than the inflation volatility in the first 1000 periods. This is because a timid conduct of the monetary policy can have perverse volatility effects as explained in section 5.3. At time 2400 a drastic monetary policy action cuts the β^* down to about 0.7. This change is sufficient to prevent multiplicity and able to reduce inflation volatility at the MSV level.

6.2 More-than-two islands

The generalization of the model to cases with more than two islands is quite straightforward even if analytically cumbersome. In this section I show two simulations with four and eight symmetrical types of professional forecasters. The only modification is that now the professional forecasters receive a signal about the average expectation of all the others. Figures 8 and 9 make clear that similar qualitative results occur. Figure 8 is generated with the same parametrization and shock series as figure 6 but with a slightly bigger gain ($g = 0.0125/\sigma$). The high volatility regime originates after 500 periods and a few periods after time 2400 a sufficiently more aggressive monetary policy makes the MSV the unique equilibrium, and so globally stable. The extent of the excess non-fundamental volatility turns out to be only slightly smaller with respect the previous exercises.

[figure 8 about here]

¹²A series of equally likely endogenous switches from the MSV to high REX and vice versa can be obtained for a parametrization that make the two basins of attraction have almost the same size.

The two jumps in the inflation dynamics in figure 9 are both endogenous. They are obtained with the same parametrization as figure 8 but with a much smaller gain $g = 0.005/\sigma$ and a smaller correlation $\rho = 0.2$. As the number of islands increases the high REX values decrease. This relation is intuitive as in the limit one expects the standard case in which REX collapse to the MSV.

[figure 9 about here]

All the same, the dynamics with more than two islands are much more sensitive to changes in correlation among observational errors. In both the above examples a decrease of 0.1 in ρ is enough to prevent the first endogenous unlucky jump. This comes from the fact that decentralized coordination with more islands is in principle more demanding. Nonetheless, increasing the number of islands also has the effect of shrinking both basins of attraction and so jumps are more likely.

7 Conclusion

This study has identified a rational for reconciling the "good luck or good policy" explanations of macro-volatility switches. The concept of rational exuberance equilibria was introduced in a simple monetary model where agents have heterogeneous expectations and, in particular, they are segmented into a few types according to the signals they receive regarding other agents' expectations of inflation.

On the one hand, equilibria entailing non-fundamental volatility can occur when agents put weight on the expectational signals as predictors of business fluctuations. Moreover, when expectations are adaptively formed using constant-gain algorithms, self-fulfilling rational exuberance can arise endogenously as the economy jumps from a fundamental equilibrium, where expectational signals are ignored, to a rational exuberance equilibrium, where agents put weight on expectational signals.

On the other hand, a decrease in non-fundamental volatility can be (even if not necessarily) the result of a drastic tightening of inflation-targeting policy that reduces the impact of expectations in the economy and prevents a multiplicity of equilibria. Nevertheless, timid "improvements" in monetary conduct are counter-productive when exuberance is already in play. The discontinuity in the effect of monetary policy could partly illuminate the difficulties in the implementation of a gradual recovery towards full stability and, on the contrary, the success of a big policy change such as that made possible by the arrival of a new governor.

The paper has focused on the importance of uncertainty about others' expectations as one of the possible sources of non-fundamental volatility. It pointed out the role of monetary policy in affecting the impact of expectations on the real economy beyond fundamentals, and provided a new argument in favor of tight inflation-targeting.

Appendix A

Households

An index $j \in (0, 1)$ denotes a continuum of representative households and $i \in (0, 1)$ a continuum of different goods. Each household j consumes a basket of all the goods produced in the economy, and supplies a type j labor input specific to the production of the good $i = j$. Households type j solve

$$\max_{\{C_{j,t}; N_{j,t}; B_{j,t}\}} \sum_{t=0}^{\infty} \beta^t \left(\log C_{j,t} - \frac{N_{j,t}^{1+\zeta}}{1+\zeta} \right),$$

subject to

$$P_t C_{j,t} + P_t B_{j,t} = P_t W_{j,t} N_{j,t} + \int \Pi_{i,t} di + (1 + r_{t-1}) P_{t-1} B_{j,t-1},$$

where $B_{j,t}$ is a bond stock, $W_{j,t}$ is the real wage, $\int \Pi_{i,t}$ is profit deriving from equally distributed ownership of firms¹³, r_t is the rate of interest,

$$C_{j,t} \equiv \left(\int C_{j,i,t}^{\frac{\theta-1}{\theta}} di \right)^{\frac{\theta}{\theta-1}} \quad \text{and} \quad P_t \equiv \left(\int P_{i,t}^{1-\theta} di \right)^{\frac{1}{1-\theta}}$$

are CES indexes with $C_{j,i,t}$ and $P_{i,t}$ being respectively consumption by agent j of the good type i , and the price of the good type i . No-Ponzi conditions apply. Trade is frictionless. The optimal supply of labor type j is determined by

$$W_{j,t} = N_{j,t}^{\zeta} C_{j,t}, \tag{16}$$

by combining first-order conditions on labor and consumption. The individual cost-minimizing demand for good i by agent j , namely $C_{j,i,t}$, is given by $C_{j,i,t} = (P_{i,t}/P_t)^{-\theta} C_{j,t}$, where θ is the CES coefficient. The total demand function for good i is given by

$$Y_{i,t} = \int \left(\frac{P_{i,t}}{P_t} \right)^{-\theta} C_{j,t} dj = Y_t \left(\frac{P_{i,t}}{P_t} \right)^{-\theta}. \tag{17}$$

In equilibrium the market clears so that the total supply

$$Y_t = \int Y_{i,t} di = \int \int C_{j,i,t} dj di, \tag{18}$$

is equal to the total demand across agents for all different goods. This is assumed to be known by the firms, which independently decide their production as explained in the following.

¹³This assumption guarantees that wealth across agents is uniformly distributed irrespective of possibly different ex-post profitability.

Firms

Each firm produces a differentiated good and sells it in a monopolistic competitive market. Firms set their prices to maximize profits. The technology for the production of the good i is given by the following

$$Y_{i,t} = e^{\psi_{t-1}} N_{i,t}, \quad (19)$$

where ψ_t is a white noise stochastic disturbance with finite variance, and $N_{i,t}$ is quantity of labour type $j = i$ hired. Firm i solves the problem $\max_{\{P_{i,t}\}} \Pi_{i,t}$ where $\Pi_{i,t} \equiv P_{i,t} Y_{i,t} - W_{i,t} P_t N_{i,t}$. Using (19) and clearing conditions, and then substituting for (16), one can write the expression for the real marginal cost $mc_{i,t}$ as

$$mc_{i,t} = \left(e^{\psi_{t-1}} \right)^{-(\zeta+1)} Y_{i,t}^\zeta C_{i,t}.$$

Notice that $C_{i,t}$ labels consumption by agent $j = i$, and not cumulative consumption of good i . The seller's desired mark-up is determined by the usual Lerner formula as $P_{i,t} = (\theta / (\theta - 1)) mc_{i,t} P_t$ yielding the condition for optimal pricing,

$$\left(\frac{P_{i,t}}{P_t} \right)^{1+\theta\zeta} = \frac{\theta}{\theta - 1} \left(e^{\psi_{t-1}} \right)^{-(\zeta+1)} Y_{i,t}^\zeta C_{i,t} \quad (20)$$

obtained after substituting for (17) followed by some trivial manipulations. Now, let us consider a log-linear approximation of (20) around the deterministic steady state at $\bar{P}_{i,t} = \bar{P}_t = 1$ written as

$$p_{i,t} = p_t + \omega_y y_t + \omega_c c_{i,t} - \tilde{z}_{t-1},$$

where $\omega_y = \zeta / (1 + \theta\zeta)$, $\omega_c = 1 / (1 + \theta\zeta)$ and $\tilde{z}_{t-1} = ((\zeta + 1) / (1 + \theta\zeta)) \psi_{t-1}$. Lower case denotes log deviations from the steady state, that is $x_t \equiv \log X_t - \log \bar{X}$. The latter corresponds to (1) in the main text in the case of perfectly informed producers.

Appendix B

Existence of REX

Equilibria are given by the system of equations:

$$\begin{aligned}\hat{\mathbf{a}}' &= \boldsymbol{\alpha} + \beta \hat{\mathbf{a}}' \\ \hat{b} &= \frac{(\beta^*/2) \hat{c}(\hat{c} + \rho)}{(1 - (\beta^*/2)(1 - \rho))\hat{c}^2 + ((2 - \beta^*)\rho - \beta^*/2)\hat{c} + (1 - (\beta^*/2))} \\ \hat{c} &= \frac{(\beta^*/2) \hat{b}(\hat{b} + \rho)}{(1 - (\beta^*/2)(1 - \rho))\hat{b}^2 + ((2 - \beta^*)\rho - \beta^*/2)\hat{b} + (1 - (\beta^*/2))}\end{aligned}$$

remembering that $bc \neq 1$. It is easily proved by substitution that the fundamental rational expectation solution is always a rest point of the T-map. Other equilibria values are in correspondence with $\hat{b} = \hat{c}$ and are obtained as solutions to

$$\hat{c}(\hat{c}^2((1 - \beta^*/2) - (\beta^*/2)\rho) - (\beta^* - (2 - \beta^*)\rho)\hat{c} + (1 - \beta^*/2) - (\beta^*/2)\rho) = 0$$

featuring respectively the high REX values (b_+, c_+) and the low REX values (b_-, c_-) .

Computing the Jacobian of the T-map

The partial derivatives of the Jacobian of the T-map are given by

$$\begin{aligned}\frac{dT_b(b, c)}{db} &= \frac{\beta^* (1 + c)(1 + c\rho) + c(c + \rho)}{2(1 + c^2 + 2c\rho)}, \\ \frac{dT_b(b, c)}{dc} &= \frac{\beta^* 2bc^2\rho^2 + c^2\rho - bc^2 + 2bc\rho + 2c + \rho + b}{2(c^2 + 2\rho c + 1)^2}, \\ \frac{dT_c(b, c)}{db} &= \frac{\beta^* 2cb^2\rho^2 + b^2\rho - cb^2 + 2bc\rho + 2b + \rho + c}{2(b^2 + 2\rho b + 1)^2}, \\ \frac{dT_c(b, c)}{dc} &= \frac{\beta^* (1 + b)(1 + b\rho) + b(b + \rho)}{2(1 + b^2 + 2b\rho)}.\end{aligned}$$

To analyze the learnability of equilibria we have to investigate the sign of the eigenvalues of the matrix $K \equiv JT - I$ (where I is the identity matrix) evaluated at the equilibrium values $\hat{\mathbf{a}}$ and $\hat{c} = \hat{b}$. The equilibrium values $(\hat{\mathbf{a}}, \hat{b}, \hat{c})$ are learnable if and only if the matrix $K_{(\hat{\mathbf{a}}, \hat{b}, \hat{c})}$ has all negative eigenvalues. The matrix $K \equiv JT - I$ at the equilibrium values $\hat{\mathbf{a}}$ and $\hat{c} = \hat{b}$ is written as

$$K_{(\hat{\mathbf{a}}, \hat{b}, \hat{c})} = \begin{pmatrix} \beta - 1 & 0 & 0 & 0 \\ 0 & \beta - 1 & 0 & 0 \\ 0 & 0 & \left[\frac{dT_b(b, c)}{db} \right]_{(\hat{b}, \hat{c})} - 1 & \left[\frac{dT_b(b, c)}{dc} \right]_{(\hat{b}, \hat{c})} \\ 0 & 0 & \left[\frac{dT_c(b, c)}{db} \right]_{(\hat{b}, \hat{c})} & \left[\frac{dT_c(b, c)}{dc} \right]_{(\hat{b}, \hat{c})} - 1 \end{pmatrix}, \quad (22)$$

with

$$\begin{aligned} \left[\frac{dT_b(b, c)}{db} \right]_{(\hat{b}, \hat{c})} - 1 &= \frac{((\beta^*/2)(1 + \rho) - 1)\hat{b}^2 + ((\beta^*/2)(1 + 2\rho) - 2\rho)\hat{b} + (\beta^*/2) - 1}{1 + \hat{b}^2 + 2\hat{b}\rho}, \\ \left[\frac{dT_b(b, c)}{dc} \right]_{(\hat{b}, \hat{c})} &= \frac{\beta^* (2\rho - 1)\hat{b}^3 + 3\rho\hat{b}^2 + 3\hat{b} + \rho}{2 (1 + \hat{b}^2 + 2\hat{b}\rho)^2}, \\ \left[\frac{dT_c(b, c)}{dc} \right]_{(\hat{b}, \hat{c})} &= \left[\frac{dT_b(b, c)}{db} \right]_{(\hat{b}, \hat{c})} \quad \text{and} \quad \left[\frac{dT_c(b, c)}{db} \right]_{(\hat{b}, \hat{c})} = \left[\frac{dT_b(b, c)}{dc} \right]_{(\hat{b}, \hat{c})}. \end{aligned}$$

References

- [1] Adam, K., 2009. Monetary Policy and Aggregate Volatility. *Journal of Monetary Economics*, 56 s1, 1-18.
- [2] Angeletos G.M., and Werning I., 2006. Crises and Prices: Information Aggregation, Multiplicity and Volatility. *American Economic Review*, 96, 1720-1736.
- [3] Benhabib, J., Farmer R., 1999. Indeterminacy and Sunspots in Macroeconomics, in: J. B. Taylor and M. Woodford (Eds.) *Handbook of Macroeconomics*, 1A, Elsevier Science B.V.
- [4] Branch W.A., Evans G.W., 2007. Endogenous Volatility and Model Uncertainty. *Review of Economic Dynamics*, 10, 207-237.
- [5] Branch W.A., Carlson, J., Evans, G. W., McGough, B., 2009. Monetary policy, endogenous inattention, and the output-price volatility trade-off. *Economic Journal*, 119, 123-157.
- [6] Carroll C.D., 2003. Macroeconomic Expectations Of Households And Professional Forecasters. *Quarterly Journal of Economics*, 118, 269-298.
- [7] Evans G.W., Honkapohja S., 2001. *Learning and Expectations in Macroeconomics*. Princeton, Princeton University Press.
- [8] Evans G.W., Honkapohja S., Williams N., 2005. *Generalized Stochastic Gradient Learning*. CESifo Working Paper No. 1576.
- [9] Evans G.W., McGough B., 2011. Representations and Sunspot Stability. *Macroeconomic Dynamics*, 15, 80-92.
- [10] Gaballo G., 2012. *Private uncertainty and Multiplicity*. Working Paper 387, Banque de France.
- [11] Galí J., Gambetti L., 2009. On the Sources of the Great Moderation. *American Economic Journal: Macroeconomics*, 1, 26-57.
- [12] Guesnerie R., 2001. *Assessing Rational Expectations*, MIT press.
- [13] Hellwig C., Mukherji A., Tsyvinski A., 2006. Self-Fulfilling Currency Crises: The Role of Interest Rates. *American Economic Review*, 96, 1769-1787.

- [14] Kydland F.E., Prescott E.C., 1977. Rules Rather than Discretion: The Inconsistency of Optimal Plans. *Journal of Political Economy*, 85, 473–491.
- [15] Lucas R., 1973. Some International Evidence on Output-Inflation Tradeoffs. *American Economic Review*, 63, 326-34.
- [16] Marcet A., Sargent T., 1989. Convergence of Least-Square Learning in Environments with Hidden State Variables and Private Information. *The Journal of Political Economy*, 97, 1306-1322.
- [17] Mankiw N.G., Reis R., Wolfers J., 2004. Disagreement about Inflation Expectations. *NBER Macroeconomics Annual 2003*, 18, 209-270.
- [18] McConnell M.M., Perez-Quiros G., 2000. Output Fluctuations in the United States: What Has Changed Since the Early 1980's. *American Economic Review*, 40, 1464-1476.
- [19] McGough B., 2003. Statistical Learning with Time-Varying Parameters. *Macroeconomic Dynamics*, 7, 119-139 .
- [20] Orphanides A., Williams J. C., 2005. The decline of activist stabilization policy: Natural rate misperceptions, learning, and expectations. *Journal of Economic Dynamics and Control*, 29, 1927–1950.
- [21] Stock J.H., Watson M.W., 2002. Has the business Cycle Changed and Why? NBER Working paper No. 9127.
- [22] Sargent T., Williams N., Zha T., 2008. The conquest of South American Inflation. *Journal of Political Economy*, 117, pp. 211-256.
- [23] Woodford M., 2003. *Interest and Prices*, Princeton Univ. Press, Princeton.

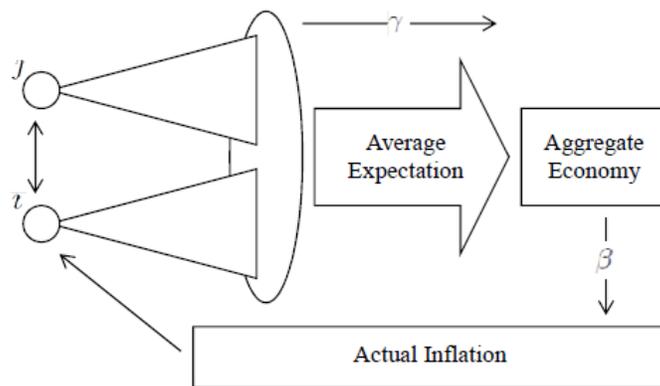


Figure 1: Information flow in the economy.

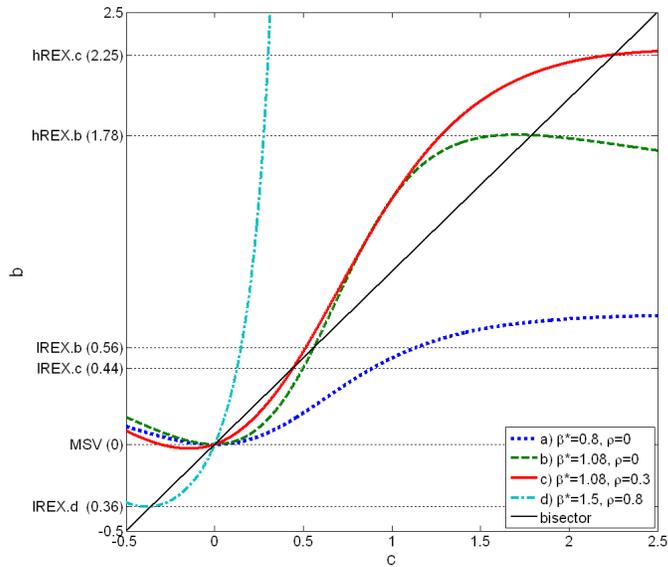


Figure 2: T-map representation for different calibrations. Equilibria lie at the intersections of the T-map with the bisector.

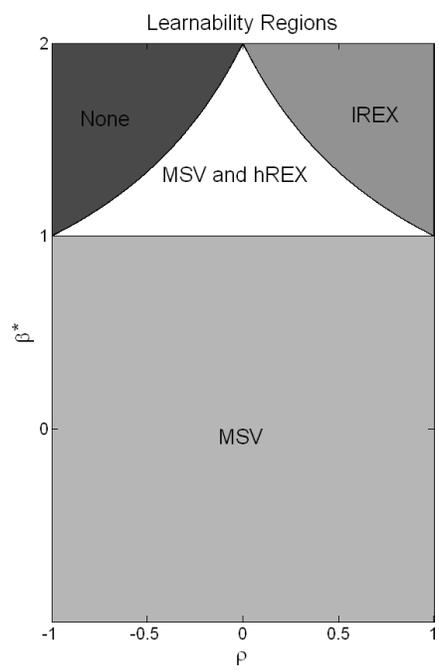


Figure 3: Numerical learnability analysis in the whole parameter space.

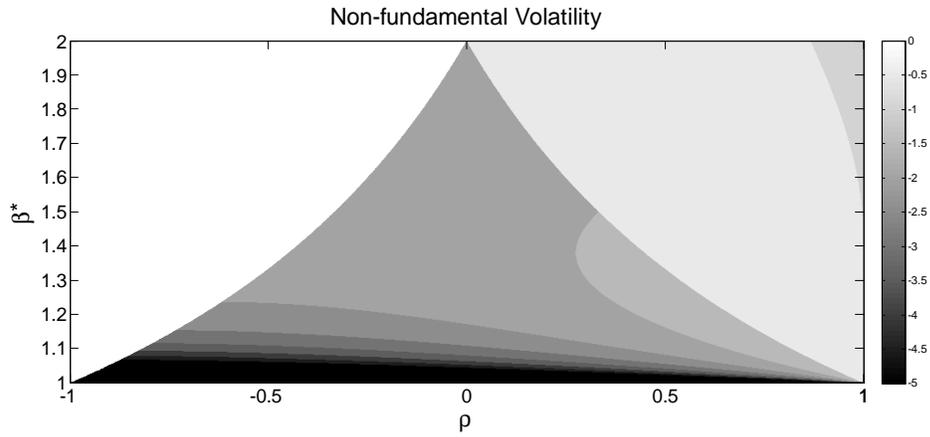


Figure 4: Numerical analysis of non-fundamental volatility. The picture shows the volatility of the aggregate expectation obtained for β^* values for which learnable REX arise. The variance of the observational errors is the unit of the measurement.

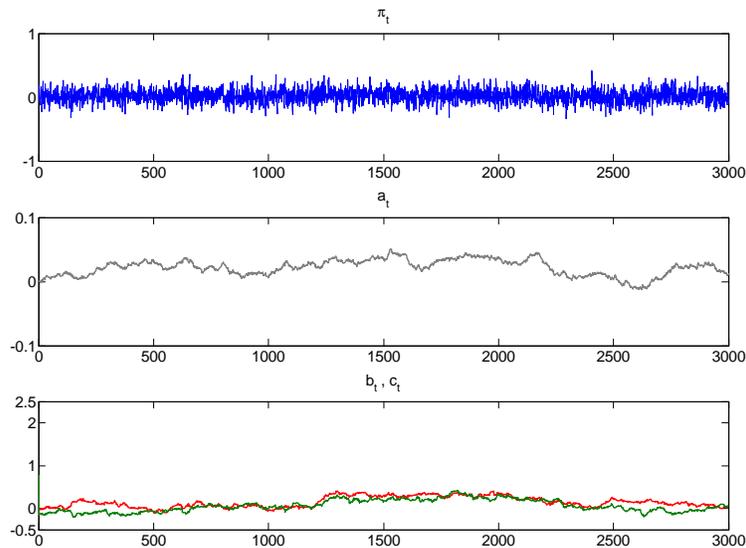


Figure 5: Benchmark case ($\beta = 0.94$, $\gamma = 0$, $\rho = 0$).

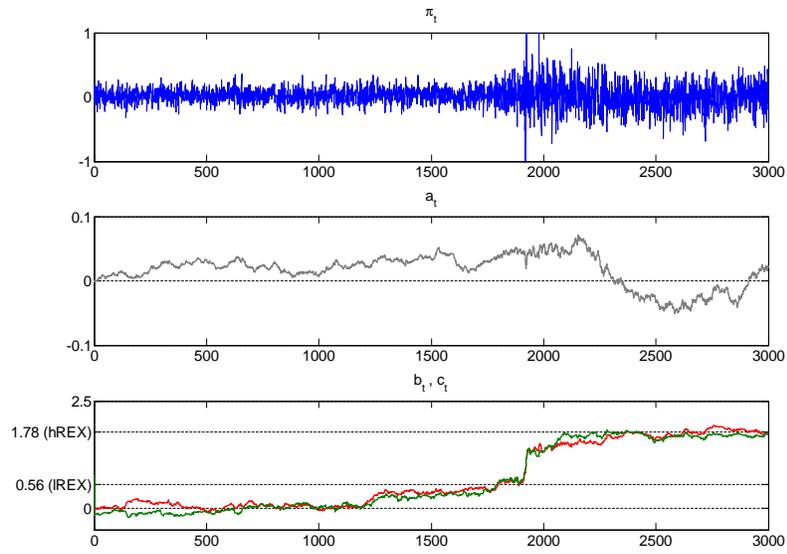


Figure 6: A case of bad luck ($\beta = 0.94$, $\gamma = 0.14$, $\rho = 0$): curve b) in figure 2.

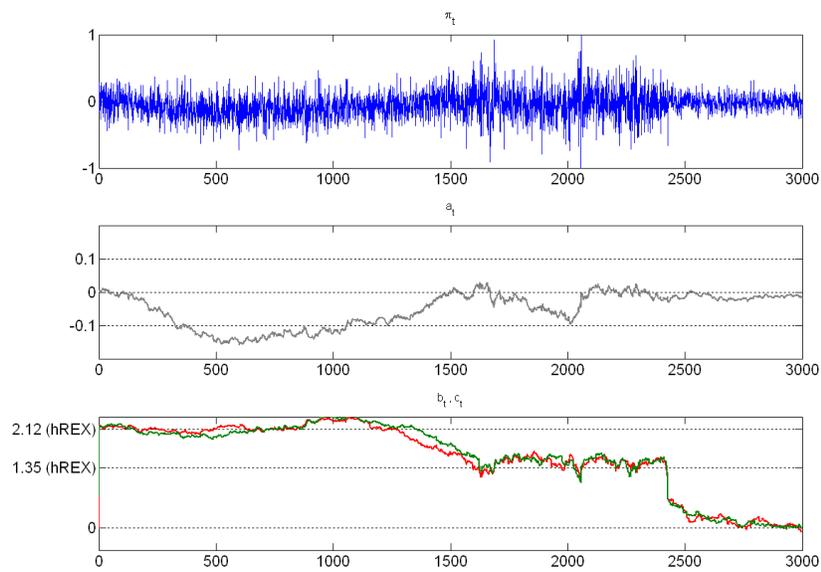


Figure 7: The conduct of monetary policy ($\beta = 0.94$, $\gamma = 0.23$, $\rho = 0$). At time 1000 a timid change in the monetary policy makes β^* fall to 1.01. At time 2400 a drastic change in the monetary policy makes β^* fall to 0.7.

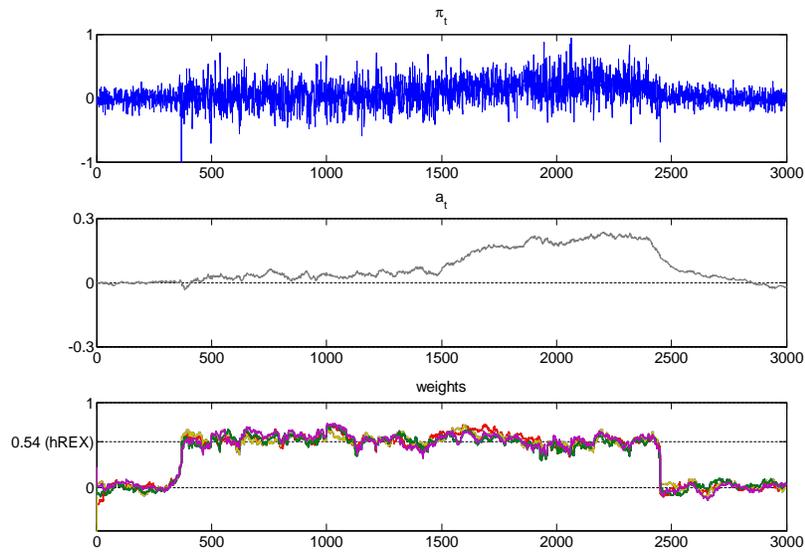


Figure 8: Good policy after bad luck with four islands ($\beta = 0.94$, $\gamma = 0.14$, $\rho = 0.3$). At time 2400 a drastic change in monetary policy makes β falls around 0.5.

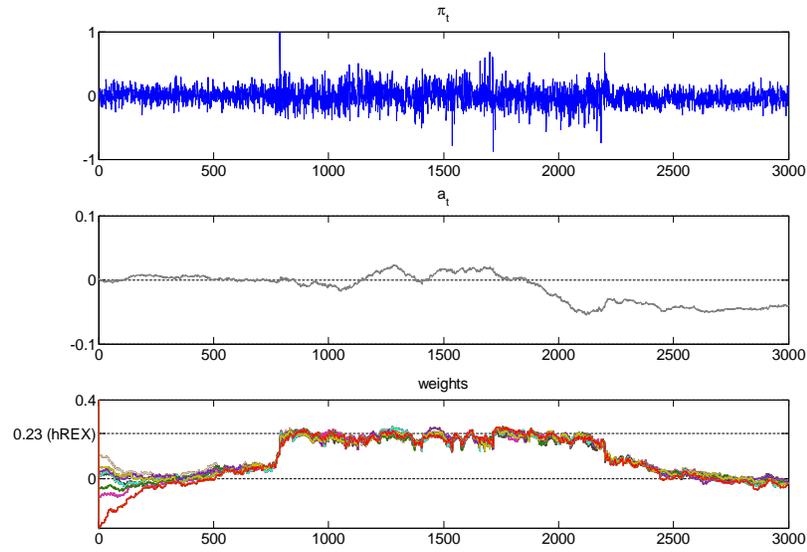


Figure 9: Good luck after bad luck with eight islands ($\beta = 0.94$, $\gamma = 0.14$, $\rho = 0.2$).

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